

EGT3  
ENGINEERING TRIPOS PART IIB

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Thursday 25 April 2024 2 to 3.40

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**Module 4M24**

**COMPUTATIONAL STATISTICS AND MACHINE LEARNING**

*Answer not more than **three** questions.*

*All questions carry the same number of marks.*

*The **approximate** percentage of marks allocated to each part of a question is indicated in the right margin.*

*Write your candidate number **not** your name on the cover sheet.*

**STATIONERY REQUIREMENTS**

Write on single-sided paper

**SPECIAL REQUIREMENTS TO BE SUPPLIED FOR THIS EXAM**

CUED approved calculator allowed

Engineering Data Book

**10 minutes reading time is allowed for this paper at the start of the exam.**

**You may not start to read the questions printed on the subsequent pages of this question paper until instructed to do so.**

**You may not remove any stationery from the Examination Room.**

1 The Monte Carlo estimator is required for the integral of the function  $f(x)$

$$\mathcal{I} = \int_a^b f(x)p(x)dx,$$

where  $x \in [a, b]$  and  $p(x)$  is the Lebesgue density. This estimator can be written as

$$\hat{\mathcal{I}} = \frac{1}{N} \sum_{n=1}^N f(x_n),$$

where each  $x_n$  is i.i.d. from  $p(x)$ .

(a) Assuming that  $\mathbb{E}\{f(X)\} = \mathcal{I}$ , and  $\text{Var}\{f(X)\} = \sigma_f^2$ , use the moment generating function to show that as  $N \rightarrow \infty$ ,  $\hat{\mathcal{I}}$  is normally distributed as

$$\hat{\mathcal{I}} \sim \mathcal{N}\left(\mathcal{I}, \frac{\sigma_f^2}{N}\right)$$

where the moment generating function for each  $f(x)$  is  $M_f(t) = \mathbb{E}\{\exp(tf(X))\}$ . [60%]

(b) The generalised Monte Carlo estimator

$$\tilde{\mathcal{I}} = \sum_{n=1}^N w_n f(x_n)$$

is unbiased when  $\sum_{n=1}^N w_n = 1$ . What choice of weights  $w_n, n = 1, \dots, N$  gives the minimum variance unbiased estimator, i.e. when  $\text{Var}\{\tilde{\mathcal{I}}\}$  is minimised? [40%]

2 The probability of data  $y \in \mathbb{R}$  given model parameters  $\theta$  is denoted by  $p(y|\theta)$ . Assuming a prior on  $\theta$  as  $p(\theta)$  the posterior follows as

$$p(\theta|y) = \frac{p(y|\theta)p(\theta)}{p(y)},$$

where

$$p(y) = \int p(y|\theta)p(\theta)d\theta.$$

Consider the introduction of a tempering or annealing variable  $t \in [0, 1]$  and  $p(t)$  is uniform in  $[0, 1]$ . The tempered data probability is  $p(y|\theta, t) \propto p(y|\theta)^t$ .

(a) Give the expression for the tempered posterior  $p(\theta|y, t)$ . [10%]

(b) Give the expression for the conditional probability  $p(y|t)$ , and its value when  $t = 0$  and  $t = 1$ , i.e.  $p(y|t = 0)$ ,  $p(y|t = 1)$ . [15%]

(c) Show that

$$\frac{d}{dt} \log p(y|t) = \mathbb{E}_{\theta|y,t} \{\log p(y|\theta)\},$$

where  $\mathbb{E}_{\theta|y,t} \{\cdot\}$  denotes the expectation with respect to the tempered posterior  $p(\theta|y, t)$ .

Hint: Note that  $x^t$  can be written in the form  $x^t = \exp(t \log(x))$ . [30%]

(d) Show that

$$\log p(y) = \mathbb{E}_{\theta,t|y} \{\log p(y|\theta)\},$$

where  $\mathbb{E}_{\theta,t|y} \{\cdot\}$  denotes the expectation with respect to the tempered joint density

$p(\theta, t|y) \propto p(\theta|y, t)p(t)$ . [15%]

(e) Derive an MCMC algorithm that gives a Monte Carlo estimate of  $\log p(y)$ . [30%]

3 The Lebesgue measure for the intervals  $[a_d, b_d]$  in  $d = 1, \dots, D$  takes the form

$$\prod_{d=1}^D (b_d - a_d).$$

Note that the Lebesgue measure is finite, monotonic and translation invariant.

(a) Show that in a complete infinite dimensional Hilbert space, the Lebesgue measure is not well defined. [35%]

(b) The standard Gaussian measure in  $\mathbb{R}$  is given as

$$g(B) = \frac{1}{\sqrt{2\pi}} \int_B \exp\left(-\frac{x^2}{2}\right) dx.$$

Consider  $\mathbb{R}^\infty$  and the product measure  $\mu = \prod_{k=1}^\infty g_k$ , with each  $g_k$  equivalent to  $g$ . Under what conditions will this product measure be well defined? What subspace of  $\mathbb{R}^\infty$  do these conditions represent? [15%]

(c) The covariance of the Gaussian measure  $\mu$  can be considered as an Identity operator. A more general linear covariance operator  $C$  on the Hilbert space  $\mathcal{H}$  can also be defined.

(i) Give conditions on  $C$  which need to be satisfied for the infinite dimensional measure  $\mu_C = \mathcal{N}(0, C)$  to be well defined. [20%]

(ii) Show that the trace operator of  $C$  in Hilbert space  $\mathcal{H}$ , having an orthonormal basis  $e_n, n = 1, \dots, \infty$ , defined on a domain  $\Omega$  can be written as

$$\sum_{n=1}^{\infty} \int_{\Omega} \int_{\Omega} e_n(x) c(x, y) e_n(y) dx dy,$$

where  $c(x, y)$  is the covariance function corresponding to the operator  $C$ . [30%]

4 Consider two univariate Gaussian probability densities on  $x \in \mathbb{R}$

$$f(x) = \frac{1}{\sqrt{2\pi\sigma_f^2}} \exp\left(-\frac{(x - \mu_f)^2}{2\sigma_f^2}\right), \quad g(x) = \frac{1}{\sqrt{2\pi\sigma_g^2}} \exp\left(-\frac{(x - \mu_g)^2}{2\sigma_g^2}\right).$$

(a) Show that the product of the two densities takes the form of a scaled (unnormalised) Gaussian  $p_{fg}(x) \propto f(x)g(x)$  with mean and variance  $\mu_{fg}, \sigma_{fg}^2$  respectively. Give the expressions for  $\mu_{fg}$  and  $\sigma_{fg}^2$ . [40%]

(b) Derive the form of the normalised density

$$p_{fg}(x) = \frac{f(x)g(x)}{\int_{\mathbb{R}} f(y)g(y)dy}.$$

[20%]

(c) Derive an Unadjusted Langevin Algorithm (ULA) to sample from  $p_{fg}(x)$  for the case where  $\mu_f = 0$ ,  $\mu_g = 1$  and  $\sigma_f = \sigma_g = 1$ . [40%]

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