# **Engineering Tripos Part IIB**

Module 4M24: Computational Statistics and Machine Learning

## 4M24 Tripos 2021/22 - Cribs

### 1. Control Variates

(a)

$$\operatorname{Var}(\hat{\mu}) = \operatorname{Var}\left(\frac{1}{N}\sum_{n=1}^{N}X_{n}\right)$$
$$= \frac{1}{N^{2}}\sum_{n=1}^{N}\operatorname{Var}(X_{n}) = \frac{1}{N^{2}}\sum_{n=1}^{N}\sigma^{2}$$
$$= \frac{\sigma^{2}}{N}$$

 $20~{\rm marks}$  available

(b) (i)

$$\hat{\mu}_{CV} = \hat{\mu} + c \left[ m - \hat{m} \right]$$

Unbiased because  $\mathbb{E} [\hat{\mu}] = \mu$ , and  $\mathbb{E} [\hat{m}] = m (= \mathbb{E} [g(x)])$ . 10 marks available

(ii)

$$\operatorname{Var}\left(\hat{\mu}_{CV}\right) = \operatorname{Var}\left(\hat{\mu} + c\left[m - \hat{m}\right]\right)$$

Using  $\operatorname{Var}(aX + bY) = a^2 \operatorname{Var}(X) + b^2 \operatorname{Var}(Y) + 2ab \operatorname{Cov}(X, Y)$  we have the following:

$$\operatorname{Var}(\hat{\mu}_{CV}) = \operatorname{Var}(\hat{\mu}) + c^{2}\operatorname{Var}(m - \hat{m}) + 2c\operatorname{Cov}(\hat{\mu}, m - \hat{m})$$
$$= \operatorname{Var}(\hat{\mu}) + c^{2}\operatorname{Var}(\hat{m}) + 2c\operatorname{Cov}(\hat{\mu}, \hat{m})$$

Now we can find the optimal variance by finding the stationary point of the variance wrt the coefficient c.

$$\frac{\partial}{\partial c} \operatorname{Var}\left(\hat{\mu}_{CV}\right) = 2c \operatorname{Var}\left(\hat{m}\right) + 2 \operatorname{Cov}\left(\hat{\mu}, \hat{m}\right) = 0$$

Solving for c:

$$c = -\frac{\operatorname{Cov}\left(\hat{\mu}, \hat{m}\right)}{\operatorname{Var}\left(\hat{m}\right)}$$

Substituting back into the expression for the control variate estimator variance  $\operatorname{Var}(\hat{\mu}) + c^2 \operatorname{Var}(\hat{m}) + 2c \operatorname{Cov}(\hat{\mu}, \hat{m})$  gives the answer:

$$\operatorname{Var}\left(\hat{\mu}_{CV}\right) = \operatorname{Var}\left(\hat{\mu}\right) - \frac{\operatorname{Cov}\left(\hat{\mu}, \hat{m}\right)^{2}}{\operatorname{Var}\left(\hat{m}\right)}$$

30 marks available

(c)  $X \sim \mathcal{U}[0,1] \in \mathbb{R}$ . First we can find the mean  $\mu = \mathbb{E}[f(x)]$ .

$$\mu = \mathbb{E}\left[\frac{1}{1+x}\right] = \int_0^1 \frac{1}{1+x} dx = [\log(1+x)]_0^1 = \log 2$$

Now to find the variance  $\sigma^2 = \mathbb{E}\left[\frac{1}{(1+x)^2}\right] - \mu^2$ .

$$\sigma^{2} = \mathbb{E}\left[\frac{1}{(1+x)^{2}}\right] - (\log 2)^{2}$$
$$= \int_{0}^{1} \frac{1}{(1+x)^{2}} dx - (\log 2)^{2}$$
$$= \left[-\frac{1}{1+x}\right]_{0}^{1} - (\log 2)^{2}$$
$$= \frac{1}{2} - (\log 2)^{2} \approx 0.0195$$

which gives

$$\operatorname{Var}(\hat{\mu}) = \frac{\sigma^2}{N} \approx \frac{0.0195}{N}$$

Defining  $\operatorname{Var}(\hat{m}) = \frac{\sigma_g^2}{N}$ . We find the mean and variance of g(x):

$$m = \mathbb{E}\left[1+x\right] = \int_0^1 (1+x)dx = \left[x + \frac{x^2}{2}\right]_0^1 = \frac{3}{2}$$
$$\sigma_g^2 = \mathbb{E}\left[(1+x)^2\right] - \frac{3}{2} = \int_0^1 (1+x)^2 dx - \frac{9}{4} = \int_0^1 (1+2x+x^2)dx - \frac{9}{4}$$
$$= \left[x + x^2 + \frac{x^3}{3}\right]_0^1 - \frac{9}{4} = \frac{1}{12}$$

which gives

$$\operatorname{Var}(\hat{m}) = \frac{1}{12N}$$

Finally we must find the covariance  $\operatorname{Cov}(\hat{\mu}, \hat{m})$ :

$$Cov(\hat{\mu}, \hat{m}) = \frac{1}{N} Cov(f(x), g(x))$$
  
=  $\frac{1}{N} \mathbb{E} [f(x)g(x)] - \frac{3}{2N} \log 2$   
=  $\frac{1}{N} \int_0^1 \frac{1+x}{1+x} dx - \frac{3}{2N} \log 2$   
=  $\frac{1}{N} \int_0^1 dx - \frac{3}{2N} \log 2$   
=  $\frac{1}{N} - \frac{3}{2N} \log 2$ 

Substituting this into the optimal variance expression gives the required result.

$$\operatorname{Var}\left(\hat{\mu}_{CV}\right) = \frac{0.0195}{N} - 12N\left(\frac{1}{N} - \frac{3}{2N}\log 2\right)^2$$

40 marks available

### 2. Function Approximation

- (a) The set  $C[-\pi,\pi]$  is the continuous functions of  $[-\pi,\pi]$  and  $L^2[-\pi,\pi]$  is the set of square integrable functions which will include discontinuous functions such as the step function, which would not be in  $C[-\pi,\pi]$ . Therefore,  $C_2[-\pi,\pi]$  would be the set of  $L^2$  integrable functions which are also continuous. 10 marks available
- (b)

$$f(x) = \sum_{k=1}^{\infty} c_k \phi_k(x)$$

$$\begin{split} \|f(x)\|_{L^{2}}^{2} &= \int_{-\pi}^{\pi} \left(\sum_{i} c_{i} \phi_{i}(x)\right) \left(\sum_{j} c_{j}^{*} \phi_{j}^{*}(x)\right) dx \\ &= \sum_{i} \sum_{j} c_{i} c_{j}^{*} \int_{-\pi}^{\pi} \phi_{i}(x) \phi_{j}^{*}(x) dx \\ &= \sum_{i} \sum_{j} c_{i} c_{j}^{*} \delta_{i,j} = \sum_{i=1}^{\infty} c_{i}^{2} \end{split}$$

where the last line comes from the orthonormality of the functions  $\{\phi_k\}$ , i.e.

$$\int_{-\pi}^{\pi} \phi_i(x)\phi_j^*(x)dx = \begin{cases} 1 & i=j\\ 0 & i\neq j \end{cases} = \delta_{i,j}$$

giving

$$||f(x)||_{L^2} = \sqrt{\sum_{i=1}^{\infty} c_i^2}$$

20 marks available

(c)

$$\|f\|_{W_2^s}^2 = \|D^s f\|_{L^2}^2$$

$$f(x) = \sum_{k=1}^{\infty} c_k \exp(ikx)$$
$$Df(x) = \sum_{k=1}^{\infty} (ik)c_k \exp(ikx)$$
$$D^s f(x) = \sum_{k=1}^{\infty} (ik)^s c_k \exp(ikx)$$

$$\|D^s f\|_{L^2}^2 = \sum_k \left| (ik)^{2s} \right| c_k^2 = \sum_k k^{2s} c_k^2$$

30 marks available

$$\begin{aligned} \epsilon_N(f) &= \|f - f_N\|_{L^2}^2 = \|\sum_{k=N+1}^{\infty} c_k \phi_k(x)\|_{L^2}^2 \\ &= \sum_{k=N+1}^{\infty} c_k^2 = \sum_{k=N+1}^{\infty} c_k^2 k^{2s} \frac{1}{k^{2s}} \\ &< \frac{1}{N^{2s}} \sum_{k=N+1}^{\infty} c_k^2 k^{2s} \\ &< \frac{1}{N^{2s}} \sum_{k=1}^{\infty} c_k^2 k^{2s} = \frac{1}{N^{2s}} \|f\|_{W_2^s}^2 \end{aligned}$$

The smoother the function, the value of s will increase, and as s increases, the impact of increasing N is greater. Rate of convergence increases as the smoothness of the function increases. 40 marks available

#### 3. Gaussian Measure

(a) (i) Since the probabilities are defined wrt the Lebesgue measure, then the definition of Bayes rule will not be valid because the Lesbesgue measure in  $\mathcal{H}$  is not well defined.

#### 10 marks available

(ii) If the reference measure  $\mu^0$  is a probability measure defined in  $\mathcal{H}$ , then the Radon-Nikodym derivative

$$\frac{d\mu^y}{d\mu^0}(x) \propto P(y|x)$$

where  $\mu^y$  is the posterior measure, and P(y|x) the likelihood of data  $y \in \mathbb{R}^D$  given  $x \in \mathcal{H}$ . 10 marks available

(b) (i) This requires that

$$\mu^{0} = \prod_{k=1}^{\infty} g \propto \exp\left(-\frac{1}{2}\sum_{k=1}^{\infty} x_{k}^{2}\right)$$

be finite, and so

$$\sum_{k=1}^{\infty} x_k^2 < \infty$$

which indicates

$$x_{k=1,\ldots,\infty} \in l_2 \subset \mathbb{R}^{\infty}$$

30 marks available

(ii)

$$\mu^{y}(x) \propto P(y|x)\mu^{0}(x) = P(y|x)\mathcal{N}(x;0,C)$$

For a proposal  $v = u + \beta \zeta$ , with  $\zeta \sim \mathcal{N}(0, C)$ , i.e.  $v \sim \mathcal{N}(u, \beta^2 C)$ , the acceptance probability is given by

$$\alpha(v, u) = \min \left\{ J(v) - J(u), 1 \right\}$$

where  $J(v) = \log P(y|v) - \frac{1}{2}|C^{-\frac{1}{2}}v|^2$ .

The problem here is that  $|C^{-\frac{1}{2}}v|^2$  is unbounded and so the acceptance ratio is not well defined. 30 marks available

(iii) Alternative proposal is  $v = \sqrt{1 - \beta^2}u + \beta\zeta$ , with  $\zeta \sim \mathcal{N}(0, C)$  - the pCN proposal. The acceptance ratio becomes

$$\alpha(v, u) = \min\left\{\frac{P(y|v)}{P(y|u)}, 1\right\}$$

since  $P(y, \cdot)$  is defined on  $\mathbb{R}^D$ , the acceptance ratio is well defined for  $u, v \in \mathcal{H}$ . 20 marks available

#### 4. Langevin Diffusion

(a) Starting with the Langevin SDE:

$$dX_t = -\nabla U(X_t)dt + \sqrt{2}dB_t$$

which has invariant density  $p(x) \propto \exp(-U(X_t))$ .

From the given SDE:

$$\nabla U(X_t) = X_t - \tanh(X_t)$$

Now integrate to find  $U(X_t)$ :

$$U(x) = \int (x - \tanh(x)) dx$$
$$= \frac{x^2}{2} - \ln(\cosh(x)) + C$$

Which gives

$$p(x) \propto \exp(-U(x))$$

$$= \exp\left(-\frac{x^2}{2} + \ln(\cosh(x)) - C\right)$$

$$\propto \cosh(x) \exp\left(-\frac{x^2}{2}\right)$$

$$= \frac{e^x + e^{-x}}{2} \exp\left(-\frac{x^2}{2}\right)$$

$$\propto \exp\left(-\frac{x^2}{2} + x\right) + \exp\left(-\frac{x^2}{2} - x\right)$$

$$\propto \exp\left(-\frac{1}{2}(x-1)^2\right) + \exp\left(-\frac{1}{2}(x+1)^2\right)$$

and so the normalised density p(x) is given by

$$p(x) = \frac{1}{2}\mathcal{N}(1,1) + \frac{1}{2}\mathcal{N}(-1,1)$$

60 marks available

(b)

$$q(x) = \exp\left(-x - \exp(-x)\right)$$

Need to determine  $-\nabla U(x) = \nabla \ln q(x)$ 

$$\ln q(x) = -x - e^{-x}$$

$$\nabla \ln q(x) = -1 + e^{-x}$$

Hence

$$-\nabla U(x) = -1 + e^{-x}$$

and the corresponding SDE is given by

$$dX_t = \left(e^{-X_t} - 1\right)dt + \sqrt{2}dB_t$$

 $40~{\rm marks}$  available