EGT2
ENGINEERING TRIPOS PART IIA

Wednesday 26 April 20179.30 to 11

Module 3F3

## SIGNAL AND PATTERN PROCESSING

Answer not more than three questions.
All questions carry the same number of marks.
The approximate percentage of marks allocated to each part of a question is indicated in the right margin.

Write your candidate number not your name on the cover sheet.

## STATIONERY REQUIREMENTS

Single-sided script paper

## SPECIAL REQUIREMENTS TO BE SUPPLIED FOR THIS EXAM

CUED approved calculator allowed
Engineering Data Book

10 minutes reading time is allowed for this paper.
You may not start to read the questions printed on the subsequent pages of this question paper until instructed to do so.

## Version SJG/4

1 (a) A source generates a stream of symbols $S_{n}, n=0,1, \ldots$ and each symbol takes one of two possible values, either A or B. The probability of symbol $S_{n}$ depends only upon the value of symbol $S_{n-1}$. Let $p\left(i_{n} \mid i_{n-1}\right)$ denote the probability that $S_{n}=i_{n}$ given $S_{n-1}=i_{n-1}$. These probabilities are given in the following table.

| $p\left(i_{n} \mid i_{n-1}\right)$ |  |  |
| :---: | :---: | :---: |
|  | $S_{n-1}=A$ | $S_{n-1}=B$ |
| $S_{n}=A$ | 0.7 | 0.1 |
| $S_{n}=B$ | 0.3 | 0.9 |

(i) Let $p\left(i_{0}\right)$ denote the probability that $S_{0}=i_{0}$. Explain how $p\left(i_{n}\right)$, the probability that $S_{n}=i_{n}$, may be calculated.
(ii) Show that a possible probability mass function for $p\left(i_{n}\right)$, for all $n \geq 0$, is

|  | $p\left(i_{n}\right)$ |
| :--- | :---: |
| $i_{n}=A$ | 0.25 |
| $i_{n}=B$ | 0.75 |

(iii) Show that the random process $S_{0}, S_{2}, S_{4}, \ldots$, generated by the same source but retaining only source symbols with even time indices, is a Markov chain, and determine its transition probability matrix.
(b) The characteristic function of a random variable $X$ is defined using the mathematical expectation $\mathbb{E}$ as $\varphi_{X}(t)=\mathbb{E}[\exp (i X t)]$ where $t$ is a real number.
(i) Let $X$ have probability density function $f_{X}(x)$. Determine the relationship between $\varphi_{X}(t)$ and the Fourier transform of $f_{X}(x)$.
(ii) Let $f_{X}(x)$ be the following triangular shaped function

$$
f_{X}(x)=1 / b\left(1-\frac{|x|}{b}\right) \quad \text { for }|x| \leq b
$$

and $f_{X}(x)=0$ for $|x|>b$. Determine $\varphi_{X}(t)$ (using the Data Book).
(iii) Express $\varphi_{X}(t)$ as a power series in $t$ and hence find $\mathbb{E}\left[X^{0}\right], \mathbb{E}\left[X^{2}\right]$ and $\mathbb{E}\left[X^{4}\right]$.

## Version SJG/4

2 A stationary, ergodic random process $\left\{X_{n}\right\}$ is measured over a time interval $n=$ $0,1, \ldots, N-1$, leading to a measured vector of samples:

$$
\mathbf{x}=\left[\begin{array}{llll}
x_{0} & x_{1} & \ldots & x_{N-1}
\end{array}\right]^{T}
$$

It is possible to estimate an unknown quantity $\theta$ using the process $\left\{X_{n}\right\}$. Let $\hat{\theta}(\mathbf{x})$ denote the estimator of $\theta$, which is a function of the measured data.
(a) Define the terms bias and variance for the estimator $\hat{\theta}$.
(b) The mean and autocorrelation function of the process are to be estimated according to the formulae:

$$
\hat{\mu}=\frac{1}{N} \sum_{n=0}^{N-1} x_{n}
$$

and

$$
\hat{R}_{X X}[k]=\frac{1}{N-k} \sum_{n=0}^{N-1-k} x_{n} x_{n+k},(k=0, \ldots, N-1) .
$$

Explain why these estimation formulae are valid, given the stated assumptions about the process.
(c) Determine whether each estimator in part (b) is unbiased.
(d) The mean value of the process is now assumed to be zero. Some autocorrelation function values are now estimated according to the above estimation formula, leading to:

$$
\hat{R}_{X X}[0]=10.5, \hat{R}_{X X}[1]=-9.1, \hat{R}_{X X}[2]=7 .
$$

It is required to predict the next value of the signal based upon previous values using a linear filter:

$$
\hat{x}_{n+1}=h_{0} x_{n}+h_{1} x_{n-1}
$$

Assuming that the estimated autocorrelation values are accurate, determine the coefficients $h_{0}$ and $h_{1}$ such that the mean-squared prediction error $\mathbb{E}\left[\left(\hat{x}_{n+1}-x_{n+1}\right)^{2}\right]$ is minimised.
(e) Determine the mean-squared prediction error of this optimal filter and compare it with a filter which takes the previous value of the process as the prediction, i.e. $\hat{x}_{n+1}=x_{n}$, commenting on why this simpler estimator would not be expected to perform well.

## Version SJG/4

3 A pilot tone in an RF communications channel is measured at the receiver in the following form:

$$
X_{n}=A+B \sin (\omega n)+V_{n}
$$

where $\left\{V_{n}\right\}$ is a white Gaussian noise process with zero mean and variance $\sigma_{V}^{2}, \omega<\pi$ is a known frequency of transmission, $A$ is an unknown DC offset and $B$ is an unknown received signal amplitude. It is required to estimate $A$ and $B$ from a measured vector of samples from the process $\left\{X_{n}\right\}$,

$$
\mathbf{x}=\left[\begin{array}{llll}
x_{0} & x_{1} & \ldots & x_{N-1}
\end{array}\right]^{T} .
$$

(a) For a particular set of parameter values $A=a$ and $B=b$, show that the total squared error term $\varepsilon=\sum_{n=0}^{N-1}\left(x_{n}-a-b \sin (\omega n)\right)^{2}$ can be expressed in terms of the unknown parameter vector $\theta=\left[\begin{array}{l}a \\ b\end{array}\right]$ as

$$
\varepsilon=\mathbf{x}^{T} \mathbf{x}-2 \mathbf{x}^{T} \mathbf{G} \theta+\theta^{T} \mathbf{G}^{T} \mathbf{G} \theta
$$

where $\mathbf{G}$ should be carefully defined.
(b) Show that the Maximum Likelihood (ML) estimator for the parameter vector can be found by minimising the following expression:

$$
N \log \left(2 \pi \sigma_{V}^{2}\right)+\frac{\varepsilon}{\sigma_{V}^{2}}
$$

and hence that the ML estimator is

$$
\theta^{\mathrm{ML}}=\mathbf{M}^{-1} \mathbf{b}
$$

where

$$
\mathbf{M}=\left[\begin{array}{cc}
N & \frac{\sin (N \omega / 2)}{\sin (\omega / 2)} \sin (\omega(N-1) / 2) \\
\frac{\sin (N \omega / 2)}{\sin (\omega / 2)} \sin (\omega(N-1) / 2) & N / 2-\frac{\sin (N \omega)}{2 \sin (\omega)} \cos (\omega(N-1))
\end{array}\right], \quad \mathbf{b}=\left[\begin{array}{c}
\sum_{n=0}^{N-1} x_{n} \\
\sum_{n=0}^{N-1} \sin (\omega n) x_{n}
\end{array}\right] .
$$

You may find the following result helpful:

$$
\sum_{n=0}^{N-1} \exp (i n b)=\exp (i(N-1) b / 2) \frac{\sin (N b / 2)}{\sin (b / 2)}
$$

## Version SJG/4

(c) With $\omega=\pi / 5$ and $N=1000$, show that the ML solution simplifies to:

$$
\hat{a}=\frac{1}{N} \sum_{n=0}^{N-1} x_{n}, \quad \hat{b}=\frac{2}{N} \sum_{n=0}^{N-1} \sin (\omega n) x_{n}
$$

(d) Explain why this simplification occurs in terms of the columns of the matrix $\mathbf{G}$. How should the data length be chosen relative to $\omega$, to ensure that this is the case?

## Version SJG/4

4 (a) For a random process $X_{n}, n=0,1, \ldots$ explain the difference between strictsense stationary (SSS) and wide-sense stationary (WSS). Why might WSS be the more practical assumption for modelling of a real-world physical process?
(b) A zero-mean random process $X_{n}$ has autocorrelation function $R_{X X}(k)=c$ for $k=0$ and $R_{X X}(k)=0$ for $|k|>0$, where $c$ is a constant. It is passed through a linear system with infinite impulse response $\left\{h_{n}\right\}_{n=-\infty}^{\infty}$. If $Y_{n}$ denotes the output process, find an expression for the autocorrelation function $R_{Y Y}(k)$ of the system output in terms of the impulse response and $c$.
(c) Find $R_{Y Y}(k)$ when $h_{n}=0$ for $n<0$ and

$$
h_{n}=a \exp (-n b)
$$

for $n \geq 0$, where $a, b$ are positive constants.
(d) Determine the power spectral density of the random process $Y_{n}$.

## END OF PAPER

